# Supplementary Materials for "The Importance of Breaking Even: How Local and Aggregate Returns 

 Make Politically Feasible Policies"
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## A Supplementary Tables for Experiment 1

Table A1: Effect of District and City-Wide Returns on Incumbent Evaluations with Returns as Categorical Variables

|  | Incumbent <br> Evaluations <br> $(-1$ to 1$)$ |
| :--- | :---: |
| District Payoff Categories: $-\$ 25$ per capita | 0.112 |
|  | $(0.092)$ |
| District Payoff Categories: $-\$ 10$ per capita | 0.130 |
|  | $(0.089)$ |
| District Payoff Categories: $-\$ 5$ per capita | $0.223^{*}$ |
|  | $(0.086)$ |
| District Payoff Categories: $-\$ 1$ per capita | $0.239^{* *}$ |
|  | $(0.081)$ |
| District Payoff Categories: $\$ 0$ per capita | $0.419^{* * *}$ |
|  | $(0.082)$ |
| District Payoff Categories: $\$ 1$ per capita | $0.458^{* * *}$ |
|  | $(0.079)$ |
| District Payoff Categories: $\$ 5$ per capita | $0.493^{* * *}$ |
|  | $(0.083)$ |
| District Payoff Categories: $\$ 10$ per capita | $0.564^{* * *}$ |
| District Payoff Categories: $\$ 25$ per capita | $(0.083)$ |
|  | $0.527^{* * *}$ |
| District Payoff Categories: $\$ 50$ per capita | $(0.085)$ |
|  | $0.610^{* * *}$ |
| City Payoff Categories: $-\$ 40$ per capita | $0.081)$ |
| City Payoff Categories: $-\$ 30$ per capita | $(0.128)$ |
|  | 0.059 |
| City Payoff Categories: $-\$ 25$ per capita | $(0.124)$ |
| City Payoff Categories: $-\$ 20$ per capita | 0.148 |
|  | $(0.119)$ |
| City Payoff Categories: $-\$ 15$ per capita | 0.122 |
|  | $(0.111)$ |
| City Payoff Categories: $-\$ 10$ per capita | 0.070 |
|  | $(0.108)$ |
|  | 0.159 |
|  | $(0.106)$ |
|  | 0.113 |
|  | $(0.106)$ |

Continued on next page

|  | Incumbent Evaluations ( -1 to 1 ) |
| :---: | :---: |
| City Payoff Categories: - \$1 per capita | $\begin{aligned} & 0.329^{* *} \\ & (0.116) \end{aligned}$ |
| City Payoff Categories: - $\$ 0.1$ per capita | $\begin{aligned} & 0.258^{*} \\ & (0.113) \end{aligned}$ |
| City Payoff Categories: $\$ 0$ per capita | $\begin{gathered} 0.388^{* * *} \\ (0.108) \end{gathered}$ |
| City Payoff Categories: $\$ 0.1$ per capita | $\begin{aligned} & 0.327^{* *} \\ & (0.110) \end{aligned}$ |
| City Payoff Categories: $\$ 1$ per capita | $\begin{gathered} 0.347^{* * *} \\ (0.102) \end{gathered}$ |
| City Payoff Categories: $\$ 5$ per capita | $\begin{gathered} 0.504^{* * *} \\ (0.108) \end{gathered}$ |
| City Payoff Categories: $\$ 10$ per capita | $\begin{gathered} 0.560^{* * *} \\ (0.109) \end{gathered}$ |
| City Payoff Categories: $\$ 15$ per capita | $\begin{gathered} 0.563^{* * *} \\ (0.106) \end{gathered}$ |
| City Payoff Categories: $\$ 20$ per capita | $\begin{gathered} 0.507^{* * *} \\ (0.106) \end{gathered}$ |
| City Payoff Categories: $\$ 25$ per capita | $\begin{aligned} & 0.342^{* *} \\ & (0.108) \end{aligned}$ |
| City Payoff Categories: $\$ 30$ per capita | $\begin{gathered} 0.515^{* * *} \\ (0.114) \end{gathered}$ |
| City Payoff Categories: $\$ 40$ per capita | $\begin{gathered} 0.534^{* * *} \\ (0.106) \end{gathered}$ |
| City Payoff Categories: $\$ 50$ per capita | $\begin{gathered} 0.591^{* * *} \\ (0.105) \end{gathered}$ |
| Vignette 2 | $\begin{gathered} -0.118^{* * *} \\ (0.031) \end{gathered}$ |
| Vignette 3 | $\begin{gathered} -0.066^{*} \\ (0.031) \end{gathered}$ |
| Constant | $\begin{gathered} -0.464^{* * *} \\ (0.110) \end{gathered}$ |
| $\mathrm{R}^{2}$ | 0.179 |
| Observations | 1487 |
| Respondents | 496 |

[^0]Table A2: Model for Simulated Electoral Tradeoffs when City Per Capita Return is Greater than 0, Experiment 1 (Table 2)

|  | Vote for Incumbent vs. Challenger (0 or 100) |
| :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{gathered} 24.257^{* * *} \\ (6.163) \end{gathered}$ |
| District Benefits (District > 0) | $\begin{gathered} 2.611 \\ (6.528) \end{gathered}$ |
| District Returns Per Capita | $\begin{gathered} 0.131 \\ (0.117) \end{gathered}$ |
| District Worse Off than City | $\begin{aligned} & -5.476 \\ & (6.213) \end{aligned}$ |
| City Returns Per Capita | $\begin{gathered} 0.069 \\ (0.121) \end{gathered}$ |
| Vignette 2 | $\begin{gathered} 1.381 \\ (4.118) \end{gathered}$ |
| Vignette 3 | $\begin{aligned} & -0.505 \\ & (3.983) \end{aligned}$ |
| Constant | $\begin{gathered} 38.324^{* * *} \\ (6.297) \end{gathered}$ |
| $\mathrm{R}^{2}$ | 0.107 |
| Observations | 711 |
| Respondents | 436 |

Table A3: Summary Statistics for Experiment 1

| Variable | Mean | SD | Min | Max |
| :--- | :---: | :---: | :---: | :---: |
| Incumbent Evaluation | 0.137 | 0.653 | -1 | 1 |
| Vote for Incumbent | 0.141 | 0.589 | -1 | 1 |
| Project Evaluation | 0.143 | 0.666 | -1 | 1 |
| District At Least Breaks Even (District $\geq 0)$ | 0.578 | 0.494 | 0 | 1 |
| District Benefits (District $>0)$ | 0.454 | 0.498 | 0 | 1 |
| City At Least Breaks Even (City $\geq 0)$ | 0.525 | 0.500 | 0 | 1 |
| City Benefits (City > 0) | 0.478 | 0.500 | 0 | 1 |
| District Worse Off than City | 0.465 | 0.499 | 0 | 1 |
| City Returns Per Capita | 0.125 | 23.628 | -50 | 50 |
| District Returns Per Capita | 0.786 | 22.563 | -50 | 50 |

Table A4: Akaike Information Criterion (AIC) by City and District Cutpoints, Experiment 1

|  | $-\$ 625 \mathrm{~K}$ | $-\$ 250 \mathrm{~K}$ | $-\$ 125 \mathrm{~K}$ | $-\$ 25 \mathrm{~K}$ | $\$ 0$ | $\$ 25 \mathrm{~K}$ | $\$ 125 \mathrm{~K}$ | $\$ 250 \mathrm{~K}$ | $\$ 625 \mathrm{~K}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $-\$ 10 \mathrm{M}$ | 2751.835 | 2741.681 | 2737.189 | 2722.278 | $\mathbf{2 7 2 1 . 9 0 6}$ | 2735.898 | 2744.931 | 2745.048 | 2745.280 |
| $-\$ 7.5 \mathrm{M}$ | 2751.655 | 2741.567 | 2737.037 | 2722.209 | $\mathbf{2 7 2 1 . 8 5 0}$ | 2735.806 | 2744.751 | 2744.936 | 2745.150 |
| $-\$ 6.25 \mathrm{M}$ | 2752.705 | 2742.270 | 2737.908 | 2722.865 | $\mathbf{2 7 2 2 . 4 8 6}$ | 2736.632 | 2745.879 | 2745.839 | 2746.214 |
| $-\$ 5 \mathrm{M}$ | 2752.637 | 2742.152 | 2737.762 | 2722.815 | $\mathbf{2 7 2 2 . 4 4 6}$ | 2736.560 | 2745.808 | 2745.716 | 2746.077 |
| $-\$ 3.75 \mathrm{M}$ | 2748.759 | 2737.883 | 2734.057 | 2719.314 | $\mathbf{2 7 1 8 . 8 9 4}$ | 2732.801 | 2741.995 | 2741.872 | 2742.297 |
| $-\$ 2.5 \mathrm{M}$ | 2747.104 | 2736.398 | 2732.290 | 2718.347 | $\mathbf{2 7 1 8 . 1 7 4}$ | 2731.859 | 2740.998 | 2740.504 | 2740.427 |
| $-\$ 1.25 \mathrm{M}$ | 2735.312 | 2725.566 | 2721.596 | 2707.728 | $\mathbf{2 7 0 7 . 3 4 1}$ | 2720.317 | 2729.395 | 2728.901 | 2729.334 |
| $-\$ 250 \mathrm{~K}$ | 2735.254 | 2725.556 | 2721.640 | 2708.084 | $\mathbf{2 7 0 7 . 7 6 3}$ | 2720.604 | 2729.575 | 2728.849 | 2729.183 |
| $-\$ 25 \mathrm{~K}$ | 2734.625 | 2725.545 | 2721.453 | 2707.851 | $\mathbf{2 7 0 7 . 5 5 5}$ | 2720.473 | 2729.444 | 2728.008 | 2728.625 |
| $\$ 0$ | 2735.162 | 2726.109 | 2721.961 | 2708.296 | $\mathbf{2 7 0 8 . 0 3 5}$ | 2721.092 | 2730.061 | 2728.405 | 2729.079 |
| $\$ 25 \mathrm{~K}$ | 2738.822 | 2730.335 | 2725.889 | 2712.104 | $\mathbf{2 7 1 1 . 8 6 6}$ | 2724.971 | 2733.637 | 2732.102 | 2732.653 |
| $\$ 250 \mathrm{~K}$ | 2740.488 | 2731.820 | 2726.898 | 2713.132 | $\mathbf{2 7 1 3 . 0 0 1}$ | 2726.334 | 2735.310 | 2734.116 | 2734.396 |
| $\$ 1.25 \mathrm{M}$ | 2742.639 | 2733.112 | 2727.880 | $\mathbf{2 7 1 3 . 6 5 5}$ | 2713.766 | 2728.155 | 2737.215 | 2736.149 | 2736.472 |
| $\$ 2.5 \mathrm{M}$ | 2745.092 | 2735.592 | 2731.236 | 2717.243 | $\mathbf{2 7 1 6 . 9 3 0}$ | 2730.205 | 2739.224 | 2739.032 | 2739.053 |
| $\$ 3.75 \mathrm{M}$ | 2746.741 | 2735.784 | 2731.817 | 2717.187 | $\mathbf{2 7 1 6 . 8 4 3}$ | 2730.761 | 2739.981 | 2739.686 | 2740.379 |
| $\$ 5 \mathrm{M}$ | 2745.127 | 2734.801 | 2730.202 | 2715.705 | $\mathbf{2 7 1 5 . 6 1 8}$ | 2729.880 | 2738.626 | 2738.006 | 2739.052 |
| $\$ 6.25 \mathrm{M}$ | 2744.689 | 2734.598 | 2729.866 | 2714.936 | $\mathbf{2 7 1 4 . 7 5 0}$ | 2729.226 | 2738.096 | 2737.238 | 2738.469 |
| $\$ 7.5 \mathrm{M}$ | 2750.425 | 2739.722 | 2735.356 | 2721.007 | $\mathbf{2 7 2 0 . 7 8 7}$ | 2734.730 | 2743.965 | 2743.930 | 2744.280 |
| $\$ 10 \mathrm{M}$ | 2750.542 | 2739.913 | 2735.608 | 2721.078 | $\mathbf{2 7 2 0 . 7 9 5}$ | 2734.661 | 2744.183 | 2744.005 | 2744.282 |

Note: Cells represent the AIC for each model with the given city and district cutpoint dummy variables, with the rows representing city cutpoint values and the columns representing district cutpoint values. The model estimated was Incumbent Evaluation $=\beta_{0}+\beta_{1} \times$ District Above Cutpoint $+\beta_{2} \times$ District At or Above Cutpoint $+\beta_{3} \times$ Aggregate Above Cutpoint $+\beta_{4} \times$ Aggregate At or Above Cutpoint $+\beta_{5} \times$ District Per Capita Less than City Per Capita $+\beta_{6} \times$ City Returns Per Capita $+\beta_{7} \times$ District Returns Per Capita $+\epsilon$. Bold values indicate lowest AIC (i.e. best fitting) model for each given set of city cutpoints.

Table A5: Heterogeneous Effect of District and City-Wide Returns on Incumbent Evaluations by Key Demographic Groups, Experiment 1

|  | Incumbent Evaluations ( -1 to 1 ) | Incumbent Evaluations ( -1 to 1 ) | Incumbent Evaluations ( -1 to 1 ) | Incumbent <br> Evaluations ( -1 to 1 ) | Incumbent Evaluations ( -1 to 1 ) |
| :---: | :---: | :---: | :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{gathered} \hline 0.199^{* * *} \\ (0.056) \end{gathered}$ | $\begin{gathered} \hline 0.148 \\ (0.118) \end{gathered}$ | $\begin{gathered} \hline 0.183 \\ (0.094) \end{gathered}$ | $\begin{aligned} & \hline 0.201^{* *} \\ & (0.067) \end{aligned}$ | $\begin{aligned} & \hline 0.293^{*} \\ & (0.133) \end{aligned}$ |
| District Benefits (District $>0$ ) | $\begin{gathered} 0.053 \\ (0.057) \end{gathered}$ | $\begin{gathered} 0.133 \\ (0.108) \end{gathered}$ | $\begin{gathered} 0.119 \\ (0.091) \end{gathered}$ | $\begin{gathered} 0.033 \\ (0.065) \end{gathered}$ | $\begin{aligned} & -0.022 \\ & (0.124) \end{aligned}$ |
| District Returns Per Capita | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.003^{* * *} \\ (0.001) \end{gathered}$ | $\begin{aligned} & 0.004^{* *} \\ & (0.001) \end{aligned}$ | $\begin{gathered} 0.003^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ |
| District Worse Off than City | $\begin{aligned} & -0.029 \\ & (0.053) \end{aligned}$ | $\begin{aligned} & -0.028 \\ & (0.052) \end{aligned}$ | $\begin{aligned} & -0.022 \\ & (0.052) \end{aligned}$ | $\begin{aligned} & -0.028 \\ & (0.052) \end{aligned}$ | $\begin{aligned} & -0.025 \\ & (0.052) \end{aligned}$ |
| City At Least Breaks Even ( City $\geq 0$ ) | $\begin{aligned} & 0.176^{*} \\ & (0.076) \end{aligned}$ | $\begin{aligned} & 0.182^{*} \\ & (0.076) \end{aligned}$ | $\begin{aligned} & 0.179^{*} \\ & (0.076) \end{aligned}$ | $\begin{aligned} & 0.181^{*} \\ & (0.076) \end{aligned}$ | $\begin{aligned} & 0.178^{*} \\ & (0.076) \end{aligned}$ |
| City Benefits ( City > 0) | $\begin{gathered} 0.018 \\ (0.076) \end{gathered}$ | $\begin{gathered} 0.013 \\ (0.076) \end{gathered}$ | $\begin{gathered} 0.017 \\ (0.076) \end{gathered}$ | $\begin{gathered} 0.014 \\ (0.076) \end{gathered}$ | $\begin{gathered} 0.009 \\ (0.076) \end{gathered}$ |
| City Returns Per Capita | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ |
| Party (7-Point Scale) | $\begin{aligned} & -0.006 \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.016) \end{aligned}$ |  |  |  |
| District At Least Breaks Even $\times$ Party |  | $\begin{gathered} 0.014 \\ (0.025) \end{gathered}$ |  |  |  |
| District Benefits $\times$ Party |  | $\begin{aligned} & -0.022 \\ & (0.023) \end{aligned}$ |  |  |  |
| Household Income (24-Point Scale) | $\begin{gathered} 0.005 \\ (0.004) \end{gathered}$ |  | $\begin{gathered} 0.006 \\ (0.006) \end{gathered}$ |  |  |
| District At Least Breaks Even $\times$ HHI |  |  | $\begin{gathered} 0.003 \\ (0.009) \end{gathered}$ |  |  |
| District Benefits $\times$ HHI |  |  | $\begin{aligned} & -0.008 \\ & (0.009) \end{aligned}$ |  |  |
| Non-White | $\begin{aligned} & -0.006 \\ & (0.047) \end{aligned}$ |  |  | $\begin{aligned} & -0.030 \\ & (0.064) \end{aligned}$ |  |
| District At Least Breaks Even $\times$ Non-White |  |  |  | $\begin{gathered} 0.012 \\ (0.118) \end{gathered}$ |  |
| District Benefits $\times$ Non-White |  |  |  | $\begin{gathered} 0.042 \\ (0.118) \end{gathered}$ |  |
| Education (8-Point Scale) | $\begin{aligned} & -0.002 \\ & (0.012) \end{aligned}$ |  |  |  | $\begin{gathered} 0.009 \\ (0.018) \end{gathered}$ |
| District At Least Breaks Even $\times$ Education |  |  |  |  | $\begin{aligned} & -0.021 \\ & (0.029) \end{aligned}$ |
| District Benefits $\times$ Education |  |  |  |  | $\begin{gathered} 0.016 \\ (0.027) \end{gathered}$ |
| Vignette 2 | $\begin{gathered} -0.109^{* * *} \\ (0.031) \end{gathered}$ | $\begin{gathered} -0.106^{* * *} \\ (0.030) \end{gathered}$ | $\begin{gathered} -0.116^{* * *} \\ (0.031) \end{gathered}$ | $\begin{gathered} -0.107^{* * *} \\ (0.031) \end{gathered}$ | $\begin{gathered} -0.103^{* * *} \\ (0.031) \end{gathered}$ |
| Vignette 3 | $\begin{gathered} -0.065^{*} \\ (0.030) \end{gathered}$ | $\begin{gathered} -0.061^{*} \\ (0.030) \end{gathered}$ | $\begin{gathered} -0.068^{*} \\ (0.030) \end{gathered}$ | $\begin{gathered} -0.061^{*} \\ (0.030) \end{gathered}$ | $\begin{gathered} -0.059^{*} \\ (0.030) \end{gathered}$ |
| Constant | $\begin{aligned} & -0.044 \\ & (0.081) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.032 \\ & (0.077) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.092 \\ & (0.065) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.028 \\ & (0.053) \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.075 \\ & (0.087) \\ & \hline \end{aligned}$ |
| $\mathrm{R}^{2}$ | 0.172 | 0.167 | 0.172 | 0.166 | 0.167 |
| Observations | 1455 | 1487 | 1461 | 1487 | 1479 |
| Respondents | 485 | 496 | 487 | 496 | 493 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

Table A6: Likelihood Ratio Tests, Experiment 1

|  | $\chi^{2}$ | $p$-value |
| :--- | :---: | :---: |
| City Return Treatment | 133.314 | 0.643 |
| District Return Treatment | 67.296 | 0.569 |

Note: We conduct a multinomial logistic regression of the treatment variable on covariates party identification, household income, race, education, and region of the survey taker to test if our treatments of city and district net return were associated with any key covariates. We then conducted a likelihood ratio test of this model with the null model for each treatment variable outcome. Each column reflects the likelihood ratio difference between a saturated and null model, as well as the $p$-value for rejecting the null hypothesis. Since we fail to reject the null for both variables, we are unable to say there is a significant association between our treatment variables and these covariates.

Table A7: Effect of District Returns on Incumbent Evaluations by City Outcome, Experiment 1

|  | Incumbent <br> Evaluations $(-1 \text { to } 1)$ | Incumbent <br> Evaluations with Covariates ( -1 to 1 ) |
| :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{gathered} 0.039 \\ (0.094) \end{gathered}$ | $\begin{gathered} 0.037 \\ (0.095) \end{gathered}$ |
| District Benefits (District > 0) | $\begin{gathered} 0.104 \\ (0.092) \end{gathered}$ | $\begin{gathered} 0.105 \\ (0.093) \end{gathered}$ |
| District Returns Per Capita | $\begin{gathered} 0.002 \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.002 \\ (0.001) \end{gathered}$ |
| District Worse Off than City | $\begin{aligned} & -0.008 \\ & (0.053) \end{aligned}$ | $\begin{aligned} & -0.013 \\ & (0.053) \end{aligned}$ |
| City At Least Breaks Even ( City $\geq 0$ ) | $\begin{gathered} 0.052 \\ (0.073) \end{gathered}$ | $\begin{gathered} 0.054 \\ (0.074) \end{gathered}$ |
| City Returns Per Capita | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.004^{* *} \\ (0.001) \end{gathered}$ |
| District At Least Breaks Even $\times$ City At Least Breaks Even | $\begin{aligned} & 0.305^{* *} \\ & (0.111) \end{aligned}$ | $\begin{gathered} 0.291^{*} \\ (0.113) \end{gathered}$ |
| District Benefits $\times$ City At Least Breaks Even | $\begin{gathered} -0.100 \\ (0.109) \end{gathered}$ | $\begin{aligned} & -0.073 \\ & (0.111) \end{aligned}$ |
| District Returns Per Capita $\times$ City At Least Breaks Even | $\begin{aligned} & 0.004^{*} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.004^{*} \\ & (0.002) \end{aligned}$ |
| Vignette 2 | $\begin{gathered} -0.105^{* * *} \\ (0.030) \end{gathered}$ | $\begin{gathered} -0.107^{* * *} \\ (0.031) \end{gathered}$ |
| Vignette 3 | $\begin{aligned} & -0.055 \\ & (0.030) \end{aligned}$ | $\begin{gathered} -0.061^{*} \\ (0.030) \end{gathered}$ |
| Constant | $\begin{gathered} 0.023 \\ (0.056) \end{gathered}$ | $\begin{gathered} 0.098 \\ (0.092) \end{gathered}$ |
| $\mathrm{R}^{2}$ | 0.185 | 0.195 |
| Observations | 1487 | 1455 |
| Respondents | 496 | 485 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent. Covariates included in Column 2 include party identification, household income, education level, geographic region, and race.

Table A8: Effect of Districts and City Returns when Returns are Substantial and Positive, Experiment 1

|  | Incumbent <br> Evaluations $(-1 \text { to } 1)$ |
| :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{aligned} & 0.189^{* *} \\ & (0.059) \end{aligned}$ |
| District Benefits (District $>0$ ) | $\begin{gathered} 0.045 \\ (0.058) \end{gathered}$ |
| District Returns Per Capita | $\begin{aligned} & 0.005^{* *} \\ & (0.002) \end{aligned}$ |
| District Worse Off than City | $\begin{aligned} & -0.005 \\ & (0.053) \end{aligned}$ |
| City At Least Breaks Even (City $\geq 0$ ) | $\begin{gathered} 0.185^{* * *} \\ (0.051) \end{gathered}$ |
| City Returns Per Capita | $\begin{aligned} & 0.004^{* *} \\ & (0.001) \end{aligned}$ |
| District Return is Substantial and Positive | $\begin{gathered} 0.053 \\ (0.059) \end{gathered}$ |
| Substantial District Return $\times$ City Returns PC | $\begin{gathered} 0.002 \\ (0.002) \end{gathered}$ |
| Substantial District Return $\times$ District Returns PC | $\begin{aligned} & -0.003 \\ & (0.002) \end{aligned}$ |
| Vignette 2 | $\begin{gathered} -0.105^{* * *} \\ (0.031) \end{gathered}$ |
| Vignette 3 | $\begin{aligned} & -0.060 \\ & (0.030) \end{aligned}$ |
| Constant | $\begin{aligned} & -0.028 \\ & (0.051) \end{aligned}$ |
| $\mathrm{R}^{2}$ | 0.169 |
| Observations | 1487 |
| Respondents | 496 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent. Substantial district returns are those that are greater than 10 .

Table A9: Effect of District and City Net Returns when the City Per Capita Return is Greater than the District Per Capita Return, Experiment 1

|  | Incumbent <br> Evaluations <br> $(-1$ to 1$)$ |
| :--- | :---: |
| District At Least Breaks Even (District $\geq 0)$ | $0.354^{* * *}$ |
|  | $(0.069)$ |
| District Benefits (District $>0)$ | 0.003 |
|  | $(0.067)$ |
| District Returns Per Capita | $0.005^{* *}$ |
|  | $(0.002)$ |
| City At Least Breaks Even (City $\geq 0)$ | -0.104 |
|  | $(0.144)$ |
| City Benefits (City $>0)$ | 0.098 |
|  | $(0.130)$ |
| City Returns Per Capita | $0.003^{*}$ |
|  | $(0.001)$ |
| Vignette 2 | -0.069 |
|  | $(0.049)$ |
| Vignette 3 | -0.022 |
|  | $(0.050)$ |
| Constant | 0.070 |
|  | $(0.093)$ |
| $\mathrm{R}^{2}$ | 0.194 |
| Observations | 692 |
| Respondents | 436 |

[^1]Table A10: Effect of District and City-Wide Returns on Incumbent Evaluations by Round, Experiment 1

|  | Round 1 | Round 2 | Round 3 |
| :--- | :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0)$ | 0.128 | $0.352^{* * *}$ | 0.129 |
|  | $(0.093)$ | $(0.090)$ | $(0.101)$ |
| District Benefits (District $>0)$ | -0.032 | -0.078 | $0.257^{* *}$ |
|  | $(0.096)$ | $(0.089)$ | $(0.095)$ |
| District Returns Per Capita | 0.002 | $0.004^{*}$ | $0.004^{*}$ |
|  | $(0.002)$ | $(0.002)$ | $(0.002)$ |
| District Worse Off than City | -0.172 | 0.016 | 0.062 |
|  | $(0.090)$ | $(0.089)$ | $(0.094)$ |
| City At Least Breaks Even (City $\geq 0)$ | $0.390^{* *}$ | -0.019 | $0.249^{*}$ |
|  | $(0.126)$ | $(0.138)$ | $(0.113)$ |
| City Benefits (City $>0)$ | -0.189 | $0.292^{*}$ | -0.124 |
|  | $(0.125)$ | $(0.134)$ | $(0.113)$ |
| City Returns Per Capita | $0.004^{*}$ | 0.003 | $0.006^{* *}$ |
|  | $(0.002)$ | $(0.002)$ | $(0.002)$ |
| Constant | 0.098 | $-0.227^{* *}$ | $-0.163^{*}$ |
|  | $(0.072)$ | $(0.080)$ | $(0.080)$ |
| $\mathrm{R}^{2}$ | 0.101 | 0.189 | 0.244 |
| Observations | 496 | 496 | 495 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: The dependent variable in each column is Incumbent Evaluation, coded -1 to 1 . Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

## B Supplementary Tables for Experiment R1

Table B1: Effect of District and City-Wide Returns on Evaluations, Experiment R1

|  | Incumbent <br> Evaluations <br> $(-1$ to 1$)$ | Vote for Incumbent <br> vs. Challenger <br> $(-1$ to 1$)$ | Project <br> Evaluation <br> $(-1$ to 1$)$ |
| :--- | :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0)$ | $0.123^{* * *}$ | $0.085^{* *}$ | $0.127^{* * *}$ |
| District Benefits (District $>0)$ | $(0.030)$ | $(0.029)$ | $(0.030)$ |
|  | $0.129^{* * *}$ | $0.087^{* *}$ | $0.096^{* *}$ |
| District Returns Per Capita | $(0.031)$ | $(0.029)$ | $(0.030)$ |
|  | $0.004^{* * *}$ | $0.003^{* * *}$ | $0.004^{* * *}$ |
| District Worse Off than City | $(0.001)$ | $(0.001)$ | $(0.001)$ |
|  | 0.048 | 0.017 | 0.052 |
| City At Least Breaks Even $($ City $\geq 0)$ | $(0.032)$ | $(0.030)$ | $(0.032)$ |
|  | $0.177^{* * *}$ | $0.085^{*}$ | $0.178^{* * *}$ |
| City Benefits (City $>0)$ | $(0.039)$ | $(0.037)$ | $(0.040)$ |
| City Returns Per Capita | 0.027 | $0.081^{*}$ | 0.053 |
|  | $(0.041)$ | $(0.039)$ | $(0.041)$ |
| Vignette 2 | $0.003^{* * *}$ | $0.002^{* * *}$ | $0.003^{* * *}$ |
|  | $(0.001)$ | $(0.001)$ | $(0.001)$ |
| Vignette 3 | $-0.089^{* * *}$ | $-0.062^{* * *}$ | $-0.114^{* * *}$ |
|  | $(0.019)$ | $(0.017)$ | $(0.018)$ |
| Constant | $-0.090^{* * *}$ | $-0.057^{* * *}$ | $-0.116^{* * *}$ |
|  | $(0.019)$ | $(0.017)$ | $(0.019)$ |
| R 2 | -0.013 | 0.041 | 0.007 |
| Observations | $(0.030)$ | $(0.027)$ | $(0.031)$ |

[^2]Table B2: Model for Simulated Electoral Tradeoffs when City Per Capita Return is Greater than 0, Experiment R1 (Table B3)

|  | Vote for Incumbent <br> vs. Challenger <br> $(0$ or 100$)$ |
| :--- | :---: |
| District At Least Breaks Even (District $\geq 0)$ | $12.359^{* * *}$ |
|  | $(3.582)$ |
| District Benefits (District $>0)$ | 2.061 |
|  | $(3.817)$ |
| District Returns Per Capita | $0.377^{* * *}$ |
|  | $(0.070)$ |
| District Worse Off than City | 6.323 |
|  | $(3.829)$ |
| City Returns Per Capita | -0.051 |
|  | $(0.072)$ |
| Vignette 2 | -1.628 |
|  | $(2.375)$ |
| Vignette 3 | 4.259 |
|  | $(2.343)$ |
| Constant | $42.416^{* * *}$ |
|  | $(3.914)$ |
| $\mathrm{R}^{2}$ | 0.067 |
| Observations | 2061 |
| Respondents | 1273 |

[^3]Table B3: Simulated Electoral Tradeoffs from Side Payoffs when City-Wide Per Capita Returns are \$1, Experiment R1

|  | Initial Losing Coalition (4-6) |  | New Winning Coalitions ( $6-4$ and $10-0$ ) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | 4 Winners 0 Break Even 6 Losers | (\$4 district return) (\$0 district return) ( $-\$ 1$ district return) | 4 Winners 2 Break Even 4 Losers | (\$3.5 district return) ( $\$ 0$ district return) ( $-\$ 1$ district return) | 4 Winners 6 Break Even 0 Losers | (\$2.5 district return) ( $\$ 0$ district return) (-\$1 district return) |
| Support in Winning Districts | 58.29\% (4) |  | 58.10\% (4) |  | 57.73\% (4) |  |
| Support in Break Even Districts | N/A (0) |  | 61.05\% (2) |  | 61.05\% (6) |  |
| Support in Losing Districts | 48.31\% (6) |  | 48.31\% (4) |  | N/A (0) |  |

Note: This table presents the predicted electoral outcomes estimated from Table B2. We assume a 10-district city passes a policy that has a per capita net return of $\$ 1$ to the city as a whole. By type (winning, break even, losing), all districts are assumed to have the same net return. Districts that come out behind have a per capita net return of $-\$ 1$. In to 2 losing districts so that they reach a per capita return of $\$ 0$, to create a minimum winning coalition of 6 . In column 3 , the councilors provide payoffs to 6 losing districts to create a universal coalition of 10 .

Table B4: Akaike Information Criterion (AIC) by City and District Cutpoints, Experiment R1

|  | $-\$ 625 \mathrm{~K}$ | $-\$ 250 \mathrm{~K}$ | $-\$ 125 \mathrm{~K}$ | $-\$ 25 \mathrm{~K}$ | $\$ 0$ | $\$ 25 \mathrm{~K}$ | $\$ 125 \mathrm{~K}$ | $\$ 250 \mathrm{~K}$ | $\$ 625 \mathrm{~K}$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $-\$ 10 \mathrm{M}$ | 8105.240 | 8116.591 | 8096.010 | 8056.825 | $\mathbf{8 0 4 0 . 6 9 8}$ | 8057.521 | 8096.397 | 8102.626 | 8119.096 |
| $-\$ 7.5 \mathrm{M}$ | 8104.119 | 8115.522 | 8094.964 | 8056.011 | $\mathbf{8 0 3 9 . 9 5 8}$ | 8056.775 | 8095.684 | 8101.416 | 8118.111 |
| $-\$ 6.25 \mathrm{M}$ | 8101.330 | 8112.994 | 8092.687 | 8053.797 | $\mathbf{8 0 3 7 . 9 7 0}$ | 8054.793 | 8093.281 | 8098.831 | 8115.474 |
| $-\$ 5 \mathrm{M}$ | 8103.091 | 8114.347 | 8094.001 | 8055.167 | $\mathbf{8 0 3 9 . 3 0 9}$ | 8056.115 | 8094.521 | 8100.665 | 8116.778 |
| $-\$ 3.75 \mathrm{M}$ | 8090.722 | 8101.563 | 8081.984 | 8043.677 | $\mathbf{8 0 2 7 . 6 3 2}$ | 8043.927 | 8082.301 | 8088.599 | 8104.408 |
| $-\$ 2.5 \mathrm{M}$ | 8076.476 | 8087.106 | 8067.731 | 8030.818 | $\mathbf{8 0 1 6 . 7 9 7}$ | 8033.791 | 8070.042 | 8075.907 | 8090.049 |
| $-\$ 1.25 \mathrm{M}$ | 8051.399 | 8062.418 | 8043.802 | 8009.227 | $\mathbf{7 9 9 5 . 6 4 2}$ | 8011.495 | 8046.130 | 8052.120 | 8065.574 |
| $-\$ 250 \mathrm{~K}$ | 8047.363 | 8058.266 | 8040.101 | 8006.450 | $\mathbf{7 9 9 3 . 0 2 0}$ | 8008.550 | 8042.853 | 8048.111 | 8061.427 |
| $-\$ 25 \mathrm{~K}$ | 8047.362 | 8058.896 | 8040.837 | 8007.262 | $\mathbf{7 9 9 2 . 9 0 5}$ | 8007.714 | 8042.516 | 8048.362 | 8062.343 |
| $\$ 0$ | 8051.172 | 8063.188 | 8044.700 | 8010.239 | $\mathbf{7 9 9 5 . 3 9 1}$ | 8010.216 | 8045.393 | 8052.320 | 8066.678 |
| $\$ 25 \mathrm{~K}$ | 8069.764 | 8081.580 | 8062.305 | 8026.571 | $\mathbf{8 0 1 1 . 5 6 5}$ | 8026.985 | 8062.886 | 8069.762 | 8084.268 |
| $\$ 250 \mathrm{~K}$ | 8089.853 | 8101.750 | 8081.564 | 8043.814 | $\mathbf{8 0 2 8 . 0 8 3}$ | 8044.001 | 8081.060 | 8087.893 | 8103.996 |
| $\$ 1.25 \mathrm{M}$ | 8096.409 | 8108.212 | 8087.813 | 8049.256 | $\mathbf{8 0 3 3 . 0 8 8}$ | 8049.380 | 8087.830 | 8094.349 | 8110.328 |
| $\$ 2.5 \mathrm{M}$ | 8099.704 | 8110.895 | 8090.727 | 8051.796 | $\mathbf{8 0 3 6 . 2 6 5}$ | 8053.336 | 8091.572 | 8097.467 | 8113.375 |
| $\$ 3.75 \mathrm{M}$ | 8092.928 | 8103.691 | 8084.207 | 8045.644 | $\mathbf{8 0 2 9 . 1 4 8}$ | 8045.152 | 8083.715 | 8089.625 | 8106.072 |
| $\$ 5 \mathrm{M}$ | 8093.812 | 8104.473 | 8085.035 | 8046.393 | $\mathbf{8 0 3 0 . 3 8 9}$ | 8046.907 | 8085.325 | 8091.466 | 8106.863 |
| $\$ 6.25 \mathrm{M}$ | 8088.104 | 8098.938 | 8079.173 | 8041.457 | $\mathbf{8 0 2 6 . 4 1 3}$ | 8043.082 | 8080.423 | 8085.676 | 8101.374 |
| $\$ 7.5 \mathrm{M}$ | 8088.319 | 8099.164 | 8079.252 | 8041.469 | $\mathbf{8 0 2 6 . 4 5 5}$ | 8043.192 | 8080.498 | 8085.727 | 8101.574 |
| $\$ 10 \mathrm{M}$ | 8092.700 | 8103.409 | 8082.986 | 8044.988 | $\mathbf{8 0 2 9 . 1 0 7}$ | 8045.310 | 8083.211 | 8089.615 | 8105.601 |

[^4]Table B5: Likelihood Ratio Tests, Experiment R1

|  | $\chi^{2}$ | $p$-value |
| :--- | :---: | :---: |
| City Return Treatment | 142.938 | 0.415 |
| District Return Treatment | 57.168 | 0.865 |

Note: We conduct a multinomial logistic regression of the treatment variable on covariates party identification, household income, race, education, and region of the survey taker to test if our treatments of city and district net return were associated with any key covariates. We then conducted a likelihood ratio test of this model with the null model for each treatment variable outcome. Each column reflects the likelihood ratio difference between a saturated and null model, as well as the $p$-value for rejecting the null hypothesis. Since we fail to reject the null for both variables, we are unable to say there is a significant association between our treatment variables and these covariates.

## C Supplementary Tables for Experiment R2

Table C1: Effect of District and County-Wide Returns on Evaluations, Experiment R2

|  | Incumbent <br> Evaluations ( -1 to 1 ) | Vote for Incumbent vs. Challenger ( -1 to 1 ) | Project <br> Evaluation (-1 to 1 ) |
| :---: | :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{gathered} 0.238^{* * *} \\ (0.018) \end{gathered}$ | $\begin{gathered} 0.184^{* * *} \\ (0.015) \end{gathered}$ | $\begin{gathered} 0.239^{* * *} \\ (0.017) \end{gathered}$ |
| District Benefits (District $>0$ ) | $\begin{gathered} 0.136^{* * *} \\ (0.018) \end{gathered}$ | $\begin{gathered} 0.103^{* * *} \\ (0.017) \end{gathered}$ | $\begin{gathered} 0.118^{* *} \\ (0.018) \end{gathered}$ |
| District Worse Off than County | $\begin{aligned} & -0.018 \\ & (0.022) \end{aligned}$ | $\begin{gathered} -0.039 \\ (0.020) \end{gathered}$ | $\begin{aligned} & -0.012 \\ & (0.022) \end{aligned}$ |
| District Returns Per Capita | $\begin{gathered} 0.005^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.003^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.005^{* * *} \\ (0.000) \end{gathered}$ |
| County Returns Per Capita | $\begin{aligned} & 0.001^{*} \\ & (0.000) \end{aligned}$ | $\begin{gathered} 0.001 \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.001 \\ (0.000) \end{gathered}$ |
| Vignette 2 | $\begin{gathered} -0.065^{* * *} \\ (0.018) \end{gathered}$ | $\begin{gathered} -0.036^{*} \\ (0.017) \end{gathered}$ | $\begin{gathered} -0.085^{* * *} \\ (0.018) \end{gathered}$ |
| Vignette 3 | $\begin{gathered} -0.089^{* * *} \\ (0.019) \end{gathered}$ | $\begin{gathered} -0.041^{*} \\ (0.017) \end{gathered}$ | $\begin{gathered} -0.082^{* * *} \\ (0.019) \end{gathered}$ |
| Vignette 4 | $\begin{gathered} -0.061^{* *} \\ (0.019) \end{gathered}$ | $\begin{gathered} -0.037^{*} \\ (0.017) \end{gathered}$ | $\begin{gathered} -0.079^{* * *} \\ (0.018) \end{gathered}$ |
| Vignette 5 | $\begin{gathered} -0.064^{*} \\ (0.020) \end{gathered}$ | $\begin{aligned} & -0.028 \\ & (0.018) \end{aligned}$ | $\begin{gathered} -0.091^{* * *} \\ (0.019) \end{gathered}$ |
| Constant | $\begin{gathered} 0.162^{* * *} \\ (0.026) \end{gathered}$ | $\begin{gathered} 0.109^{* * *} \\ (0.024) \end{gathered}$ | $\begin{gathered} 0.210^{* * *} \\ (0.026) \end{gathered}$ |
| $\mathrm{R}^{2}$ | 0.172 | 0.133 | 0.163 |
| Observations | 5079 | 5079 | 5079 |
| Respondents | 1018 | 1018 | 1018 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

Table C2: Effect of District and State-Wide Returns on Evaluations, Experiment R2

|  | Incumbent Evaluations ( -1 to 1 ) | Vote for Incumbent vs. Challenger ( -1 to 1 ) | Project <br> Evaluation $(-1 \text { to } 1)$ |
| :---: | :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{gathered} 0.275^{* * *} \\ (0.018) \end{gathered}$ | $\begin{gathered} 0.176^{* * *} \\ (0.016) \end{gathered}$ | $\begin{gathered} 0.273^{* * *} \\ (0.018) \end{gathered}$ |
| District Benefits (District $>0$ ) | $\begin{gathered} 0.117^{* * *} \\ (0.019) \end{gathered}$ | $\begin{aligned} & 0.092^{* * *} \\ & (0.018) \end{aligned}$ | $\begin{gathered} 0.091^{* * *} \\ (0.019) \end{gathered}$ |
| District Worse Off than State | $\begin{aligned} & -0.028 \\ & (0.024) \end{aligned}$ | $\begin{aligned} & -0.032 \\ & (0.022) \end{aligned}$ | $\begin{gathered} -0.034 \\ (0.024) \end{gathered}$ |
| District Returns Per Capita | $\begin{gathered} 0.004^{* * *} \\ (0.000) \end{gathered}$ | $\begin{aligned} & 0.003^{* * *} \\ & (0.000) \end{aligned}$ | $\begin{gathered} 0.004^{* * *} \\ (0.000) \end{gathered}$ |
| State Returns Per Capita | $\begin{aligned} & 0.001^{*} \\ & (0.000) \end{aligned}$ | $\begin{aligned} & 0.001^{*} \\ & (0.000) \end{aligned}$ | $\begin{aligned} & 0.001^{* *} \\ & (0.000) \end{aligned}$ |
| Vignette 2 | $\begin{gathered} -0.030 \\ (0.018) \end{gathered}$ | $\begin{gathered} -0.017 \\ (0.016) \end{gathered}$ | $\begin{gathered} -0.050^{* *} \\ (0.017) \end{gathered}$ |
| Vignette 3 | $\begin{gathered} -0.016 \\ (0.019) \end{gathered}$ | $\begin{gathered} -0.015 \\ (0.017) \end{gathered}$ | $\begin{gathered} -0.064^{* * *} \\ (0.019) \end{gathered}$ |
| Vignette 4 | $\begin{gathered} -0.068^{* * *} \\ (0.020) \end{gathered}$ | $\begin{gathered} -0.050^{* *} \\ (0.017) \end{gathered}$ | $\begin{gathered} -0.100^{* * *} \\ (0.019) \end{gathered}$ |
| Vignette 5 | $\begin{gathered} -0.070^{* * *} \\ (0.020) \end{gathered}$ | $\begin{gathered} -0.049^{* *} \\ (0.018) \end{gathered}$ | $\begin{gathered} -0.107^{* * *} \\ (0.020) \end{gathered}$ |
| Constant | $\begin{gathered} 0.118^{* * *} \\ (0.028) \end{gathered}$ | $\begin{gathered} 0.089^{* * *} \\ (0.025) \end{gathered}$ | $\begin{gathered} 0.193^{* * *} \\ (0.028) \end{gathered}$ |
| $\mathrm{R}^{2}$ | 0.157 | 0.111 | 0.148 |
| Observations | 4923 | 4923 | 4923 |
| Respondents | 985 | 985 | 985 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

Table C3: Effect of District and Aggregate-Wide Returns on Evaluations with State Interactions, Experiment R2

|  | Incumbent Evaluations (-1 to 1 ) | Vote for Incumbent vs. Challenger (-1 to 1 ) | Project Evaluation (-1 to 1 ) |
| :---: | :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{gathered} 0.238^{* * *} \\ (0.018) \end{gathered}$ | $\begin{gathered} \hline 0.184^{* * *} \\ (0.015) \end{gathered}$ | $\begin{gathered} \hline 0.239^{* * *} \\ (0.017) \end{gathered}$ |
| District Benefits (District $>0$ ) | $\begin{gathered} 0.136^{* * *} \\ (0.018) \end{gathered}$ | $\begin{gathered} 0.103^{* * *} \\ (0.017) \end{gathered}$ | $\begin{gathered} 0.118^{* * *} \\ (0.018) \end{gathered}$ |
| District Worse Off than Aggregate | $\begin{aligned} & -0.018 \\ & (0.022) \end{aligned}$ | $\begin{aligned} & -0.039 \\ & (0.020) \end{aligned}$ | $\begin{aligned} & -0.012 \\ & (0.022) \end{aligned}$ |
| District Returns Per Capita | $\begin{gathered} 0.005^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.003^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.005^{* * *} \\ (0.000) \end{gathered}$ |
| Aggregate Returns Per Capita | $\begin{aligned} & 0.001^{*} \\ & (0.000) \end{aligned}$ | $\begin{gathered} 0.001 \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.001 \\ (0.000) \end{gathered}$ |
| Vignette 2 | $\begin{gathered} -0.065^{* * *} \\ (0.018) \end{gathered}$ | $\begin{gathered} -0.036^{*} \\ (0.017) \end{gathered}$ | $\begin{gathered} -0.085^{* * *} \\ (0.018) \end{gathered}$ |
| Vignette 3 | $\begin{gathered} -0.089^{* * *} \\ (0.019) \end{gathered}$ | $\begin{gathered} -0.041^{*} \\ (0.017) \end{gathered}$ | $\begin{gathered} -0.082^{* * *} \\ (0.019) \end{gathered}$ |
| Vignette 4 | $\begin{gathered} -0.061^{* *} \\ (0.019) \end{gathered}$ | $\begin{gathered} -0.037^{*} \\ (0.017) \end{gathered}$ | $\begin{gathered} -0.079^{* * *} \\ (0.018) \end{gathered}$ |
| Vignette 5 | $\begin{gathered} -0.064^{* *} \\ (0.020) \end{gathered}$ | $\begin{aligned} & -0.028 \\ & (0.018) \end{aligned}$ | $\begin{gathered} -0.091^{* * *} \\ (0.019) \end{gathered}$ |
| State | $\begin{aligned} & -0.044 \\ & (0.039) \end{aligned}$ | $\begin{aligned} & -0.020 \\ & (0.034) \end{aligned}$ | $\begin{aligned} & -0.017 \\ & (0.038) \end{aligned}$ |
| State x District At Least Breaks Even | $\begin{gathered} 0.036 \\ (0.026) \end{gathered}$ | $\begin{aligned} & -0.008 \\ & (0.021) \end{aligned}$ | $\begin{gathered} 0.034 \\ (0.025) \end{gathered}$ |
| State x District Benefits | $\begin{aligned} & -0.019 \\ & (0.026) \end{aligned}$ | $\begin{aligned} & -0.011 \\ & (0.025) \end{aligned}$ | $\begin{aligned} & -0.027 \\ & (0.026) \end{aligned}$ |
| State x District Worse Off than Aggregate | $\begin{aligned} & -0.011 \\ & (0.033) \end{aligned}$ | $\begin{gathered} 0.007 \\ (0.030) \end{gathered}$ | $\begin{aligned} & -0.023 \\ & (0.032) \end{aligned}$ |
| State x District Returns Per Capita | $\begin{aligned} & -0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.001) \end{aligned}$ | $\begin{gathered} -0.001^{*} \\ (0.001) \end{gathered}$ |
| State x Aggregate Returns Per Capita | $\begin{aligned} & -0.000 \\ & (0.001) \end{aligned}$ | $\begin{gathered} 0.000 \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.000 \\ (0.001) \end{gathered}$ |
| State x Vignette 2 | $\begin{gathered} 0.036 \\ (0.026) \end{gathered}$ | $\begin{gathered} 0.019 \\ (0.023) \end{gathered}$ | $\begin{gathered} 0.035 \\ (0.025) \end{gathered}$ |
| State x Vignette 3 | $\begin{aligned} & 0.073^{* *} \\ & (0.027) \end{aligned}$ | $\begin{gathered} 0.025 \\ (0.024) \end{gathered}$ | $\begin{gathered} 0.018 \\ (0.026) \end{gathered}$ |
| State x Vignette 4 | $\begin{aligned} & -0.007 \\ & (0.027) \end{aligned}$ | $\begin{aligned} & -0.013 \\ & (0.024) \end{aligned}$ | $\begin{aligned} & -0.020 \\ & (0.026) \end{aligned}$ |
| State x Vignette 5 | $\begin{aligned} & -0.006 \\ & (0.028) \end{aligned}$ | $\begin{aligned} & -0.021 \\ & (0.025) \end{aligned}$ | $\begin{aligned} & -0.016 \\ & (0.027) \end{aligned}$ |
| Constant | $\begin{gathered} 0.162^{* * *} \\ (0.026) \end{gathered}$ | $\begin{gathered} 0.109^{* * *} \\ (0.024) \end{gathered}$ | $\begin{gathered} 0.210^{* * *} \\ (0.026) \end{gathered}$ |
| $\mathrm{R}^{2}$ | 0.165 | 0.123 | 0.156 |
| Observations | 10002 | 10002 | 10002 |
| Respondents | 2003 | 2003 | 2003 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

Table C4: Effect of District and Aggregate-Wide Returns on Evaluations with Fixed Effects by Individual as Unit (Within Person Analysis), Experiment R2

|  | Incumbent <br> Evaluations (-1 to 1 ) | Vote for Incumbent vs. Challenger ( -1 to 1 ) | Project <br> Evaluation (-1 to 1 ) |
| :---: | :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{gathered} 0.254^{* * *} \\ (0.013) \end{gathered}$ | $\begin{gathered} 0.179^{* * *} \\ (0.011) \end{gathered}$ | $\begin{gathered} 0.255^{* * *} \\ (0.012) \end{gathered}$ |
| District Benefits ( District $>0$ ) | $\begin{gathered} 0.114^{* * *} \\ (0.012) \end{gathered}$ | $\begin{gathered} 0.088^{* * *} \\ (0.012) \end{gathered}$ | $\begin{gathered} 0.095 * * * \\ (0.012) \end{gathered}$ |
| District Worse Off than Aggregate | $\begin{gathered} -0.038^{* *} \\ (0.015) \end{gathered}$ | $\begin{gathered} -0.048^{* * *} \\ (0.013) \end{gathered}$ | $\begin{gathered} -0.036^{*} \\ (0.014) \end{gathered}$ |
| District Returns Per Capita | $\begin{gathered} 0.005^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.003^{* * *} \\ (0.000) \end{gathered}$ | $\begin{gathered} 0.004^{* * *} \\ (0.000) \end{gathered}$ |
| Aggregate Returns Per Capita | $\begin{gathered} 0.001^{* * *} \\ (0.000) \end{gathered}$ | $\begin{aligned} & 0.001^{* *} \\ & (0.000) \end{aligned}$ | $\begin{gathered} 0.001^{* * *} \\ (0.000) \end{gathered}$ |
| Vignette 2 | $\begin{gathered} -0.047^{* * *} \\ (0.013) \end{gathered}$ | $\begin{gathered} -0.027^{*} \\ (0.012) \end{gathered}$ | $\begin{gathered} -0.067^{* * *} \\ (0.012) \end{gathered}$ |
| Vignette 3 | $\begin{gathered} -0.053^{* * *} \\ (0.013) \end{gathered}$ | $\begin{gathered} -0.028^{*} \\ (0.012) \end{gathered}$ | $\begin{gathered} -0.073^{* * *} \\ (0.013) \end{gathered}$ |
| Vignette 4 | $\begin{gathered} -0.064^{* *} \\ (0.014) \end{gathered}$ | $\begin{gathered} -0.044^{* * *} \\ (0.012) \end{gathered}$ | $\begin{gathered} -0.089^{* * *} \\ (0.013) \end{gathered}$ |
| Vignette 5 | $\begin{gathered} -0.067^{* * *} \\ (0.014) \end{gathered}$ | $\begin{gathered} -0.039^{* *} \\ (0.013) \end{gathered}$ | $\begin{gathered} -0.099^{* * *} \\ (0.014) \end{gathered}$ |
| Observations | 10002 | 10002 | 10002 |
| Respondents | 2003 | 2003 | 2003 |
| $\mathrm{R}^{2}$ (full model) | 0.611 | 0.572 | 0.618 |
| $\mathrm{R}^{2}$ (proj model) | 0.297 | 0.223 | 0.289 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

Table C5: Effect of District and Aggregate-Wide Returns on Evaluations Subsetted for Vignettes 1 and 2, Experiment R2

|  | Incumbent <br> Evaluations $(-1 \text { to } 1)$ | Vote for Incumbent vs. Challenger ( -1 to 1 ) | Project <br> Evaluation $(-1 \text { to } 1)$ |
| :---: | :---: | :---: | :---: |
| District At Least Breaks Even (District $\geq 0$ ) | $\begin{gathered} 0.160^{* * *} \\ (0.025) \end{gathered}$ | $\begin{gathered} 0.098^{* * *} \\ (0.021) \end{gathered}$ | $\begin{gathered} 0.167^{* * *} \\ (0.024) \end{gathered}$ |
| District Benefits (District > 0) | $\begin{aligned} & 0.068^{*} \\ & (0.028) \end{aligned}$ | $\begin{aligned} & 0.079^{* *} \\ & (0.025) \end{aligned}$ | $\begin{gathered} 0.039 \\ (0.027) \end{gathered}$ |
| District Worse Off than Aggregate | $\begin{aligned} & -0.023 \\ & (0.029) \end{aligned}$ | $\begin{aligned} & -0.005 \\ & (0.026) \end{aligned}$ | $\begin{aligned} & -0.026 \\ & (0.029) \end{aligned}$ |
| District Returns Per Capita | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.003^{* * *} \\ (0.001) \end{gathered}$ | $\begin{gathered} 0.004^{* * *} \\ (0.001) \end{gathered}$ |
| Aggregate Returns Per Capita | $\begin{aligned} & 0.001^{*} \\ & (0.001) \end{aligned}$ | $\begin{gathered} 0.001 \\ (0.000) \end{gathered}$ | $\begin{aligned} & 0.001^{* *} \\ & (0.001) \end{aligned}$ |
| Vignette 2 | $\begin{gathered} -0.043^{* * *} \\ (0.013) \end{gathered}$ | $\begin{gathered} -0.024^{*} \\ (0.012) \end{gathered}$ | $\begin{gathered} -0.063^{* * *} \\ (0.012) \end{gathered}$ |
| Constant | $\begin{gathered} 0.213^{* * *} \\ (0.030) \end{gathered}$ | $\begin{gathered} 0.131^{* * *} \\ (0.027) \end{gathered}$ | $\begin{gathered} 0.273^{* * *} \\ (0.030) \end{gathered}$ |
| $\mathrm{R}^{2}$ | 0.098 | 0.071 | 0.089 |
| Observations | 4001 | 4001 | 4001 |
| Respondents | 2003 | 2003 | 2003 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

Table C6: Akaike Information Criterion (AIC) by City and District Cutpoints, Experiment R2

|  | $\$ 5$ | $\$ 10$ | $\$ 15$ | $\$ 25$ | $\$ 50$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $-\$ 55$ | 17413.72 | 17413.26 | 17411.36 | 17410.77 | 17410.99 |
| $-\$ 50$ | 17413.28 | 17412.82 | 17410.90 | 17410.30 | 17410.52 |
| $-\$ 45$ | 17402.93 | 17402.51 | 17400.70 | 17399.96 | 17400.15 |
| $-\$ 30$ | 17394.68 | 17394.15 | 17392.08 | 17391.34 | 17391.55 |
| $-\$ 25$ | 17373.51 | 17372.89 | 17370.73 | 17369.96 | 17370.16 |
| $-\$ 20$ | 17309.08 | 17308.35 | 17306.06 | 17305.19 | 17305.39 |
| $-\$ 15$ | 17256.44 | 17255.63 | 17252.97 | 17251.91 | 17252.15 |
| $-\$ 10$ | 17212.34 | 17211.30 | 17208.60 | 17207.59 | 17207.79 |
| $-\$ 5$ | 17169.63 | 17168.53 | 17165.91 | 17165.03 | 17165.22 |
| $-\$ 2$ | 17067.73 | 17066.69 | 17064.59 | 17063.78 | 17063.90 |
| $-\$ 0.5$ | 17002.91 | 17002.17 | 17000.71 | 16999.90 | 16999.96 |
| $-\$ 0.1$ | 16934.61 | 16934.29 | 16933.15 | 16932.43 | 16932.51 |
| $\$ 0$ | 16910.04 | 16909.87 | $\mathbf{1 6 9 0 8 . 7 9}$ | 16908.54 | $\mathbf{1 6 9 0 8 . 7 2}$ |
| $\$ 0.1$ | 17150.67 | 17150.51 | 17149.55 | 17149.55 | 17149.94 |
| $\$ 0.5$ | 17280.37 | 17280.24 | 17279.24 | 17279.21 | 17279.53 |
| $\$ 2$ | 17383.28 | 17382.89 | 17381.44 | 17381.03 | 17381.20 |
| $\$ 5$ | 17388.59 | 17388.06 | 17386.25 | 17385.34 | 17385.49 |
| $\$ 10$ | 17354.31 | 17353.64 | 17352.01 | 17351.05 | 17351.14 |
| $\$ 15$ | 17329.64 | 17328.90 | 17326.95 | 17325.92 | 17326.03 |
| $\$ 20$ | 17300.33 | 17299.55 | 17297.49 | 17296.52 | 17296.65 |
| $\$ 25$ | 17282.13 | 17281.45 | 17279.36 | 17278.53 | 17278.71 |
| $\$ 30$ | 17264.07 | 17263.53 | 17261.47 | 17260.85 | 17261.08 |
| $\$ 45$ | 17276.23 | 17275.72 | 17273.65 | 17273.08 | 17273.34 |
| $\$ 50$ | 17321.57 | 17321.01 | 17318.79 | 17318.15 | 17318.40 |
| $\$ 55$ | 17366.34 | 17365.88 | 17363.57 | 17363.08 | 17363.43 |

Note: Cells represent the AIC for each model with the given aggregate and district per capita cutpoint dummy variables. Unlike previous models, the columns represent the aggregate cutpoint values and the rows represent the district cutpoint values. The model estimated was Incumbent Evaluation $=\beta_{0}+\beta_{1} \times$ District Above Cutpoint $+\beta_{2} \times$ District At or Above Cutpoint $+\beta_{3} \times$ Aggregate Above Cutpoint $+\beta_{4} \times$ Aggregate At or Above Cutpoint $+\beta_{5} \times$ District Per Capita Less than City Per Capita $+\beta_{6} \times$ City Returns Per Capita $+\beta_{7} \times$ District Returns Per Capita $+\epsilon$. Bold values indicate lowest AIC (i.e. best fitting) model for each given set of city cutpoints.

Table C7: Likelihood Ratio Tests, Experiment R2

|  | $\chi^{2}$ | $p$-value |
| :--- | :---: | :---: |
| City Return Treatment | 0.015 | 1 |
| District Return Treatment | 111.786 | 1 |

Note: We conduct a multinomial logistic regression of the treatment variable on covariates party identification, household income, race, education, and region of the survey taker to test if our treatments of city and district net return were associated with any key covariates. We then conducted a likelihood ratio test of this model with the null model for each treatment variable outcome. Each column reflects the likelihood ratio difference between a saturated and null model, as well as the $p$-value for rejecting the null hypothesis. Since we fail to reject the null for both variables, we are unable to say there is a significant association between our treatment variables and these covariates.

## D Supplementary Tables for Experiment 2

Table D1: Summary Statistics for Experiment 2

| Variable | Mean | SD | Min | Max |
| :--- | :---: | :---: | :---: | :---: |
| Incumbent Evaluation | 61.452 | 23.975 | 0.00 | 100 |
| Vote for Incumbent | 3.388 | 1.023 | 1.00 | 5 |
| Approval of Project | 3.489 | 1.133 | 1.00 | 5 |
| Project Evaluation | 62.629 | 23.831 | 0.00 | 100 |
| District At Least Breaks Even (District $\geq 0)$ | 0.669 | 0.471 | 0.00 | 1 |
| District Benefits (District $>$ 0) | 0.511 | 0.500 | 0.00 | 1 |
| District Worse Off than City | 0.662 | 0.473 | 0.00 | 1 |
| City Returns Per Capita | 1.721 | 1.698 | 0.04 | 4 |
| District Returns Per Capita | 0.684 | 1.621 | -1.00 | 4 |
| Generic | 0.282 | 0.450 | 0.00 | 1 |
| Bad for District, Not Germane | 0.187 | 0.390 | 0.00 | 1 |
| Bad for District, Germane | 0.091 | 0.288 | 0.00 | 1 |
| Fair Share, Not Germane | 0.094 | 0.291 | 0.00 | 1 |
| Fair Share, Germane | 0.185 | 0.388 | 0.00 | 1 |

Table D2: Effect of Challenger Criticisms on Evaluations when District At Least Breaks Even, Experiment 2

|  | Project <br> Evaluation <br> $(0$ to 100) | Approval of <br> Project <br> $(1$ to 5$)$ | Incumbent <br> Evaluation <br> $(0$ to 100) | Vote for Incumbent <br> vs. Challenger <br> $(1$ to 5$)$ |
| :--- | :---: | :---: | :---: | :---: |
| District Benefits (District >0) | -0.491 | -0.015 | -0.026 | -0.017 |
|  | $(1.367)$ | $(0.065)$ | $(1.381)$ | $(0.057)$ |
| District Returns Per Capita | 0.634 | 0.027 | 0.506 | 0.030 |
|  | $(0.389)$ | $(0.019)$ | $(0.385)$ | $(0.017)$ |
| District Worse Off than City | $-3.988^{*}$ | -0.114 | $-3.628^{*}$ | -0.129 |
|  | $(1.702)$ | $(0.081)$ | $(1.698)$ | $(0.072)$ |
| City Returns Per Capita | 0.449 | 0.008 | -0.046 | 0.006 |
|  | $(0.324)$ | $(0.016)$ | $(0.331)$ | $(0.014)$ |
| Second Sample | -0.543 | 0.023 | -0.280 | 0.023 |
|  | $(0.916)$ | $(0.044)$ | $(0.925)$ | $(0.040)$ |
| Generic Critique | -0.472 | 0.033 | 0.399 | 0.019 |
|  | $(1.389)$ | $(0.066)$ | $(1.386)$ | $(0.059)$ |
| Bad Deal Critique (Not Germane) | -2.724 | -0.067 | -0.752 | 0.039 |
|  | $(1.440)$ | $(0.068)$ | $(1.439)$ | $(0.059)$ |
| Fair Share Critique (Not Germane) | -2.558 | -0.079 | -1.438 | -0.105 |
|  | $(1.721)$ | $(0.081)$ | $(1.756)$ | $(0.076)$ |
| Fair Share Critique (Germane) | $-4.356^{*}$ | -0.133 | $-4.032^{*}$ | -0.124 |
| Constant | $(1.749)$ | $(0.084)$ | $(1.770)$ | $(0.076)$ |
|  | $68.149^{* * *}$ | $3.621^{* * *}$ | $66.244^{* * *}$ | $3.514^{* * *}$ |
| $\mathrm{R}^{2}$ | $(2.622)$ | $(0.123)$ | $(2.621)$ | $(0.110)$ |
| Observations | 0.019 | 0.010 | 0.018 | 0.014 |

${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

Table D3: Likelihood Ratio Tests, Experiment 2

|  | $\chi^{2}$ | $p$-value |
| :--- | :---: | :---: |
| City Return Treatment | 16.184 | 0.094 |
| District Return Treatment | 27.251 | 0.343 |

Note: We conduct a multinomial logistic regression of the treatment variable on covariates party identification, household income, race, education, and region of the survey taker to test if our treatments of city and district net return were associated with any key covariates. We then conducted a likelihood ratio test of this model with the null model for each treatment variable outcome. Each column reflects the likelihood ratio difference between a saturated and null model, as well as the $p$-value for rejecting the null hypothesis. Since we fail to reject the null for both variables, we are unable to say there is a significant association between our treatment variables and these covariates.

## E Supplementary Figures

Figure E1: Effect of District and City Returns on Incumbent Evaluations, Experiment 1


Note: Coefficients and confidence intervals are from regression model in Table A1 in the Appendix.

Figure E2: Effect of District and City Returns on Incumbent Evaluations, Experiment R1


Note: Coefficients and confidence intervals are from the same regression model as seen in Table A1 in the Appendix, except with data from Experiment R1.

Figure E3: Mean Incumbent Evaluations by City Per Capita Returns, Binned by Net District Returns, Experiment 1


Note: The gray shaded area is the expected discontinuity interval, which ranges from $-\$ 1$ to $\$ 0$ for the district per capita return. The size of the points at each coordinate reflects the sample size for the given city-district treatment pairing. City per capita returns of $-\$ 1, \$ 0$, and $\$ 1$ were oversampled.

Figure E4: Mean Incumbent Evaluations by District Per Capita Returns, Binned by Net City Returns, Experiment R1


Note: The gray shaded area is the expected discontinuity interval, which ranges from $-\$ 1$ to $\$ 0$ for the district per capita return. The size of the points at each coordinate reflects the sample size for the given city-district treatment pairing. District per capita returns of $-\$ 1, \$ 0$, and $\$ 1$ were oversampled.

Figure E5: Mean Incumbent Evaluations by District Per Capita Returns, Binned by Aggregate Jurisdiction, Experiment R2


Note: The gray shaded area is the expected discontinuity interval, which ranges from $-\$ 0.1$ to $\$ 0$ for the district per capita return. The size of the points at each coordinate reflects the sample size for the given city-district treatment pairing. District per capita returns of $-\$ 2$ to $\$ 2$ were oversampled.

## F Survey Vignettes

## F1 Experiment 1 Vignette

Suppose that the members of your city's local government were considering a major road renovation project and your representative voted "yes" in support of the project. The project will cost $\$ 100,000$ in taxes from each district. The city has 10 districts. Each district has roughly 25,000 people and the city's total population is around 250,000 . Below is a description of the project's net return for the city as a whole and your district.

Row A of the table reports the net return to all districts (the whole city), including your own district.

Row B of the table reports the net return to your district.

The net return is the economic value produced by the project, minus the program's cost in tax dollars. Negative numbers indicate that the city or your district come out behind on the project, while positive numbers indicate that the city or your district come out ahead on the project.

|  | Net Return |
| :--- | :--- |
| Row A: Net Return to the Whole City <br> (Including Your District) | $\$ 6,250,000$ |
| Row B: Net Return to Your District | $\$ 25,000$ |

Adding the gains and losses of all the districts, including your district, the city comes out $\$ 6,250,000$ ahead.

Your district comes out $\$ 25,000$ ahead.

This means that the typical person in your city is $\$ 25.00$ better off and the typical person in your district is $\$ 1.00$ better off.

This outcome can be summarized in the following figure:


## F2 Experiment 2 Vignette

To ensure you have enough time to read the following, you will be able to proceed once 20 seconds have elapsed.

Suppose that the members of your city's local government were considering a major road renovation project and your representative voted "yes" in support of the project. The project will cost $\$ 100,000$ in taxes from each district. The city has 10 districts. Each district has roughly 25,000 people and the city's total population is around 250,000 . Below is a description of the project's net return for the city as a whole and your district.

Column A of the table reports the net return to all districts (the whole city), including your own district.

Column B of the table reports the net return to your district.

The net return is the economic value produced by the project, minus the program's cost in tax dollars. Negative numbers indicate that the city or your district come out behind on the project, while positive numbers indicate that the city or your district come out ahead on the project.

|  | Column A: Net Return to the Whole <br> City (Including Your District) | Column B: Net Return <br> to Your District |
| :--- | :--- | :--- |
| Net Return | $\$ 250,000$ | $-\$ 25,000$ |

Adding the gains and losses of all the districts, the city comes out $\$ 250,000$ ahead.
Your district comes out $\$ 25,000$ behind.

This means that the typical person in your city is $\$ 1.00$ better off and the typical person in your district is $\$ 1.00$ worse off.
\{Control Treatment
Suppose that in the upcoming election for the non-partisan city council, your councilor is being challenged by someone who says the incumbent is doing a bad job.
$\{$ Control + Generic criticism $\}$
In a debate held before the election, the challenger for your district's seat made the following statement:
"Our councilor isn't doing a good job. We're not on the right track and it's time for a change."
$\{$ Control + district performance criticism $\}$
In a debate held before the election, the challenger for your district's seat made the following statement:
"Our city is in trouble. Our councilor isn't doing a good job. We're not on the right track and it's time for a change. Just look at this road project. This project is a bad deal for the district. You don't want to reelect someone who can't look out for our district's voters."
$\{$ Control + fairness performance criticism $\}$
In a debate held before the election, the challenger for your district's seat made the following statement:
"Our councilor isn't doing a good job. We're not on the right track and it's time for a change. Just look at this road project. This project doesn't give our district its fair share. Our district puts in the same amount of money as other districts, but some other district is getting a lot more. You don't want to reelect someone who can't look out for our district's voters."

## G Survey Question Wordings

## G1 Experiments 1, R1, and R2

Do you approve of the road project adopted by the city council?

1. Strongly approve
2. Somewhat approve
3. Neither approve nor disapprove
4. Somewhat disapprove
5. Strongly disapprove

Do you approve of your incumbent councilor who voted "yes" on the road project?

1. Strongly approve
2. Somewhat approve
3. Neither approve nor disapprove
4. Somewhat disapprove
5. Strongly disapprove

Suppose you were voting in the upcoming election for city councilor in your district. Who would you vote for?

1. I would certainly vote for the incumbent
2. I would probably vote for the incumbent
3. I'd be equally likely to vote for either candidate
4. I would probably vote for the challenger
5. I would certainly vote for the challenger

## G2 Experiment 2

Do you approve of the road project adopted by the city council?

1. Strongly approve
2. Somewhat approve
3. Neither approve nor disapprove
4. Somewhat disapprove
5. Strongly disapprove

Suppose you were voting in the upcoming election for city councilor in your district. Who would you vote for?

1. I would certainly vote for the incumbent
2. I would probably vote for the incumbent
3. I'd be equally likely to vote for either candidate
4. I would probably vote for the challenger
5. I would certainly vote for the challenger

On a scale of 0 to 100 , with 0 being extremely unfavorable and 100 being extremely favorable, what is your rating of your councilor (who voted for the project)?

On a scale of 0 to 100 , with 0 being extremely unfavorable and 100 being extremely favorable, what is your rating of this road project?


[^0]:    ${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
    Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

[^1]:    ${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
    Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

[^2]:    ${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
    Note: Dependent variables are listed in each column. Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

[^3]:    ${ }^{* * *} p<0.001 ;{ }^{* *} p<0.01 ;{ }^{*} p<0.05$.
    Note: Models estimated using ordinary least squares regression, with standard errors clustered by respondent.

[^4]:    Note: Cells represent the AIC for each model with the given city and district cutpoint dummy variables, with the rows representing city cutpoint values and the columns representing district cutpoint values. The model estimated was Incumbent Evaluation $=\beta_{0}+\beta_{1} \times$ District Above Cutpoint $+\beta_{2} \times$ District At or Above Cutpoint $+\beta_{3} \times$ Aggregate Above Cutpoint $+\beta_{4} \times$ Aggregate At or Above Cutpoint $+\beta_{5} \times$ District Per Capita Less than City Per Capita $+\beta_{6} \times$ City Returns Per Capita $+\beta_{7} \times$ District Returns Per Capita $+\epsilon$. Bold values indicate lowest AIC (i.e. best fitting) model for each given set of city cutpoints.

